

Statistical research of models generated by discretely observed solution to SDE's

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Introduction

The talk is divided into two parts:

The first one is devoted to research asymptotic properties of estimators for discretely observed solution to Lévy driven SDE's

$$dX_t = a_\theta(X_t)dt + dZ_t,$$

where Z is a Lévy process without diffusion component.

The following tasks are investigated:

- Estimation of the Fischer information.
- Effectiveness of the estimation method.
- One-Step estimator.
- Software implementation of the algorithms.

We consider one parameter case and multi parameters one.

Second part

We consider the statistical model generated by SDE

$$dX_t = a(\alpha; X_t)dt + \sigma(\beta; X_t)dW_t.$$

Here, W is a one dimensional Wiener process and the coefficients a and σ are differentiable functions with bounded derivatives.

- Estimation of the transition probability density.
- Construct QMLE and One-Step estimator.
- Software implementation.

Estimation of the Fisher information

We consider the Markov process given by equation

$$dX_t = a_\theta(X_t)dt + dZ_t.$$

Let Z be a Lévy process without a diffusion component that is

$$Z_t = ct + \int_0^t \int_{|u|>1} uv(ds, du) + \int_0^t \int_{|u|\leq 1} u\tilde{\nu}(ds, du),$$

where ν is a Poisson point measure with the intensity measure $ds\mu(du)$, and $\tilde{\nu}(ds, du) = \nu(ds, du) - ds\mu(du)$ is respective compensated Poisson measure.

We assume the Lévy measure μ to satisfy the following:

(i) for some $\kappa > 0$,

$$\int_{|u| \geq 1} u^{2+\kappa} \mu(du) < \infty;$$

(ii) for some $u_0 > 0$, the restriction of μ on $[-u_0, u_0]$ has a positive density $\sigma \in C^2([-u_0, 0) \cup (0, u_0])$;

(iii) there exists C_0 such that

$$|\sigma'(u)| \leq C_0 |u|^{-1} \sigma(u), \quad |\sigma''(u)| \leq C_0 u^{-2} \sigma(u), \quad |u| \in (0, u_0];$$

(iv)

$$\left(\log \frac{1}{\varepsilon} \right)^{-1} \mu(\{u : |u| \geq \varepsilon\}) \rightarrow \infty, \quad \varepsilon \rightarrow 0.$$

Assumptions about drift

Assume the following:

- (i) Let a has bounded derivatives $\partial_{x^i \theta^j}^{i+j} a$, $i \leq 3, j \leq 2$
- (ii) Derivatives $\partial_x a$, $\partial_{xx}^2 a$, $\partial_{x\theta}^2 a$, $\partial_{xxx}^3 a$, $\partial_{xx\theta}^3 a$, $\partial_{x\theta\theta}^3 a$, $\partial_{xxx\theta}^4 a$ are bounded and

$$|a_\theta(x)| + |\partial_\theta a_\theta(x)| + |\partial_{\theta\theta}^2 a_\theta(x)| \leq C(1 + |x|).$$

for all $\theta \in \Theta$, $x \in \mathbb{R}$

- (iii) for given $\theta_0 \in \Theta$ there is a neighborhood $(\theta_-, \theta_+) \subset \Theta$ of the point θ_0 , such that

$$\limsup_{|x| \rightarrow +\infty} \frac{a_\theta(x)}{x} < 0 \quad \text{uniformly w.r.t. } \theta \in (\theta_-, \theta_+).$$

Theorem 1. Representation for the transition probability density and its logarithmic derivative

- X has a transition probability density p_t^θ w.r.t. the Lebesgue measure, which has an integral representation

$$p_t^\theta(x, y) = E_x^\theta[\Xi_t I_{X_t > y}], \quad t > 0, \quad x, y \in \mathbb{R}.$$

$$\Xi_t := \delta\left(\frac{1}{DX_t}\right) = \frac{\delta(1)}{DX_t} + \frac{D^2 X_t}{(DX_t)^2}.$$

- p_t^θ has a derivative $\partial_\theta p_t^\theta(x, y) = g_t^\theta(x, y)p_t^\theta(x, y)$ with

$$g_t^\theta(x, y) = \begin{cases} E_{x,y}^{t,\theta} \Xi_t^1, & p_t^\theta(x, y) > 0, \\ 0, & \text{otherwise.} \end{cases}$$

$$\Xi_t^1 := \delta\left(\frac{\partial_\theta X_t}{DX_t}\right) = \frac{(\partial_\theta X_t)\delta(1)}{DX_t} + \frac{(\partial_\theta X_t)D^2 X_t}{(DX_t)^2} - \frac{D(\partial_\theta X_t)}{DX_t}.$$

Fischer information

The logarithm of the transition probability density has a continuous derivative w.r.t. θ on the open subset of $(0, \infty) \times \mathbb{R} \times \mathbb{R} \times \Theta$ and, on this subset, admits the integral representation

$$\partial_\theta \log p_t^\theta(x, y) = g_t^\theta(x, y) = E_{x,y}^{t,\theta} \Xi_t^1.$$

Let $x \in \mathbb{R}$, $n \in \mathbb{N}$, $0 < t_1 < \dots < t_n$ be fixed. Respective Fischer information equals to

$$I(\theta) = \sum_{k=1}^n E_x^\theta \left(g_{t_k - t_{k-1}}^\theta(X_{t_{k-1}}, X_{t_k}) \right)^2,$$

Effectiveness checking algorithm in one-parametric model is:

Step 1: Select the method of estimation and build an estimator $\hat{\theta}_n$ of the unknown parameter θ_0

Step 2: Generate N trajectories of the process X with $\hat{\theta}_n = \theta_0$ and for each of them construct a sample size n

Step 3: Calculate estimators $\hat{\theta}_n^k, k = 1, \dots, N$ and find a sample variance $s_N^2 = \frac{1}{N} \sum_{k=1}^N (\sqrt{n}(\hat{\theta}_n^k - \theta)^2)$

Step 4: Calculate n_0 and generate N trajectories of the process X with $\hat{\theta}_n = \theta_0$, for each of them calculate $\Xi_n^k(n_0), k = 1, \dots, N$

Step 5: Find a sample mean $J_n(\hat{\theta}_n, n_0) = \frac{1}{N} \sum_{k=1}^N (\Xi_n^k(n_0))^2$

Step 6: By value $\sqrt{J_n(\hat{\theta}_n, n_0)s_N^2}$ decide on the effectiveness of the method.

Simultaneous equations to find Ξ_t^1 items

Let $Y_t^1 = \partial_\theta X_t$, $Y_t^2 = DX_t$, $Y_t^3 = D\partial_\theta X_t$, $Y_t^4 = D^2X_t$.

Then $\bar{Y}_t := (Y_t^1, Y_t^2, Y_t^3, Y_t^4)$ is the solution to:

$$\begin{cases} dY_t^1 = \partial_x a_\theta(X_t) Y_t^1 dt + \partial_\theta a_\theta(X_t) dt; \\ dY_t^2 = \partial_x a_\theta(X_t) Y_t^2 dt + dDZ_t; \\ dY_t^3 = \partial_x a_\theta(X_t) Y_t^3 dt + (\partial_{x\theta} a_\theta(X_t) Y_t^2 + \partial_{xx} a_\theta(X_t) Y_t^1 Y_t^2) dt; \\ dY_t^4 = \partial_x a_\theta(X_t) Y_t^4 dt + \partial_{xx} a_\theta(X_t) (Y_t^2)^2 dt + dD^2 Z_t; \\ Y_{t_0}^i = 0, \quad i = 1, \dots, 4, \end{cases}$$

where

$$DZ_t = \int_0^t \int_{\mathbb{R}} \varrho(u) \nu(ds, du), \quad D^2Z_t = \int_0^t \int_{\mathbb{R}} \varrho(u) \varrho'(u) \nu(ds, du),$$

$$\delta(1) = - \int_0^t \int_{\mathbb{R}} \frac{(\sigma(u) \varrho(u))'}{\sigma(u)} \tilde{\nu}(ds, du) \quad \varrho(u) = \begin{cases} u^2, & |u| \leq u_1; \\ 0, & |u| \geq u_0. \end{cases}$$

Citation

Expression of Fisher's information was presented in

- Ivanenko D.O., Kulik A.M. (2015) **Malliavin calculus approach to statistical inference for Levy driven SDE's.**
Meth. and Comp. in Appl. Prob., vol. 17, p. 107–123.

The effectiveness checking algorithm has been described in

- Ivanenko D.O., Bodnarchuk S.V. (2015) **Efficiency of estimating method, in statistical models controlled by Levy noise.**
Th. of Prob. and Math. Stat., vol. 92, p. 9 – 22.

To construct the One-step estimator we need to expression of the second derivative. The following result was presented in

- Ivanenko D.O. (2014) **Second derivative of the log-likelihood in the model given by a Lévy driven SDE's.**
Visn. Kyiv Nat. Univ., arxiv org, 1410.2880, vol. 2, p. 18 –22.

Example 1.

Figure: Efficiency in one-parameter nonlinear model.

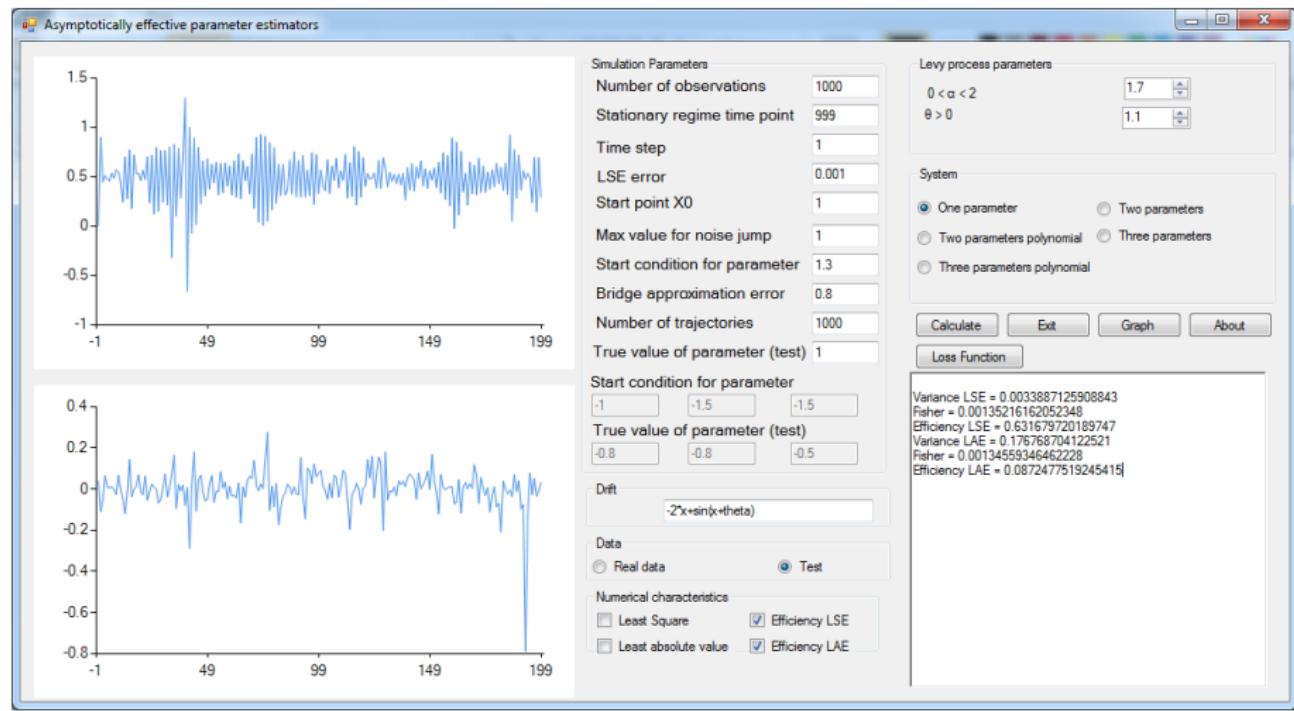
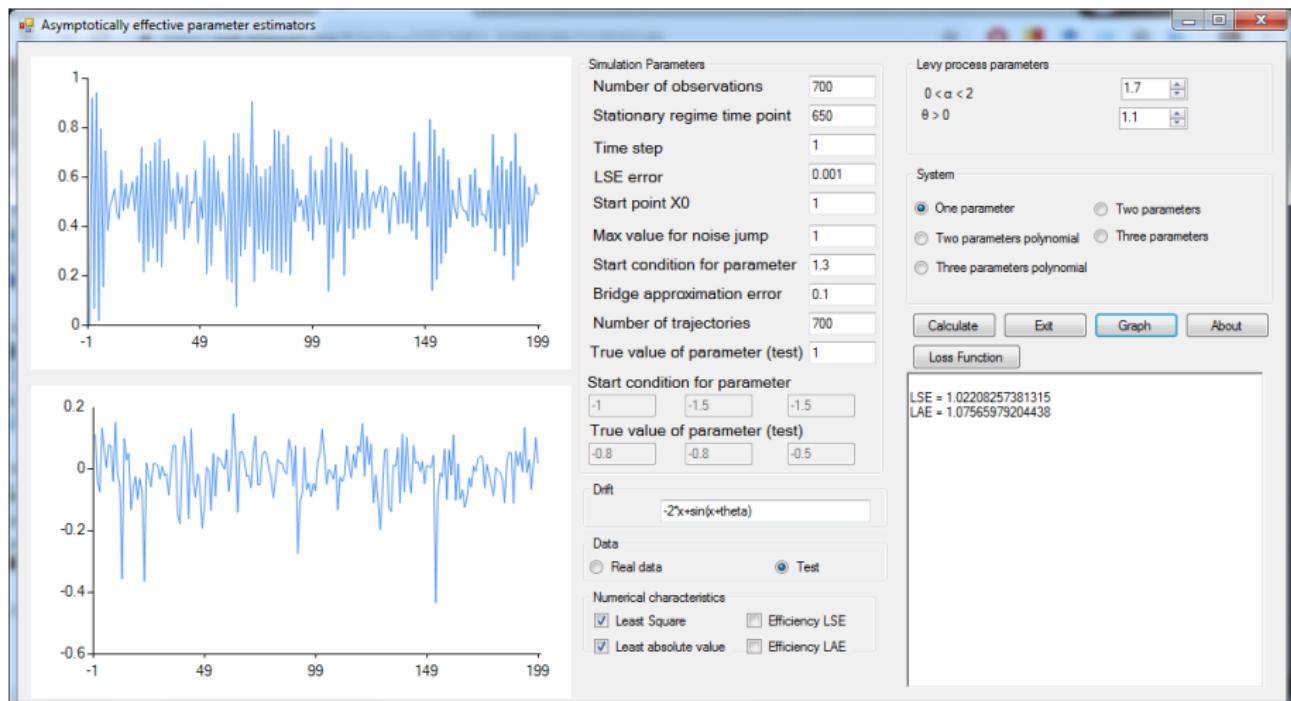


Figure: LSE and LAE in one-parameter nonlinear model.



Example 2.

Figure: Efficiency in two-parameter nonlinear model.

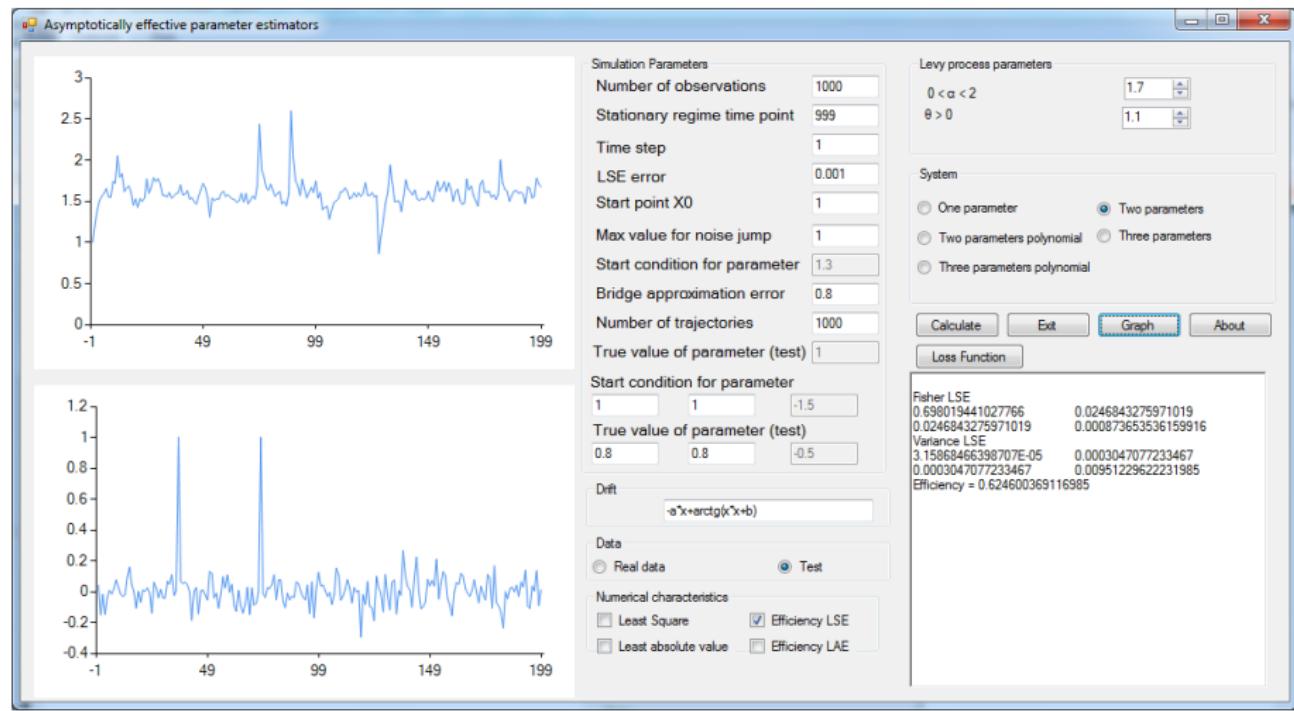
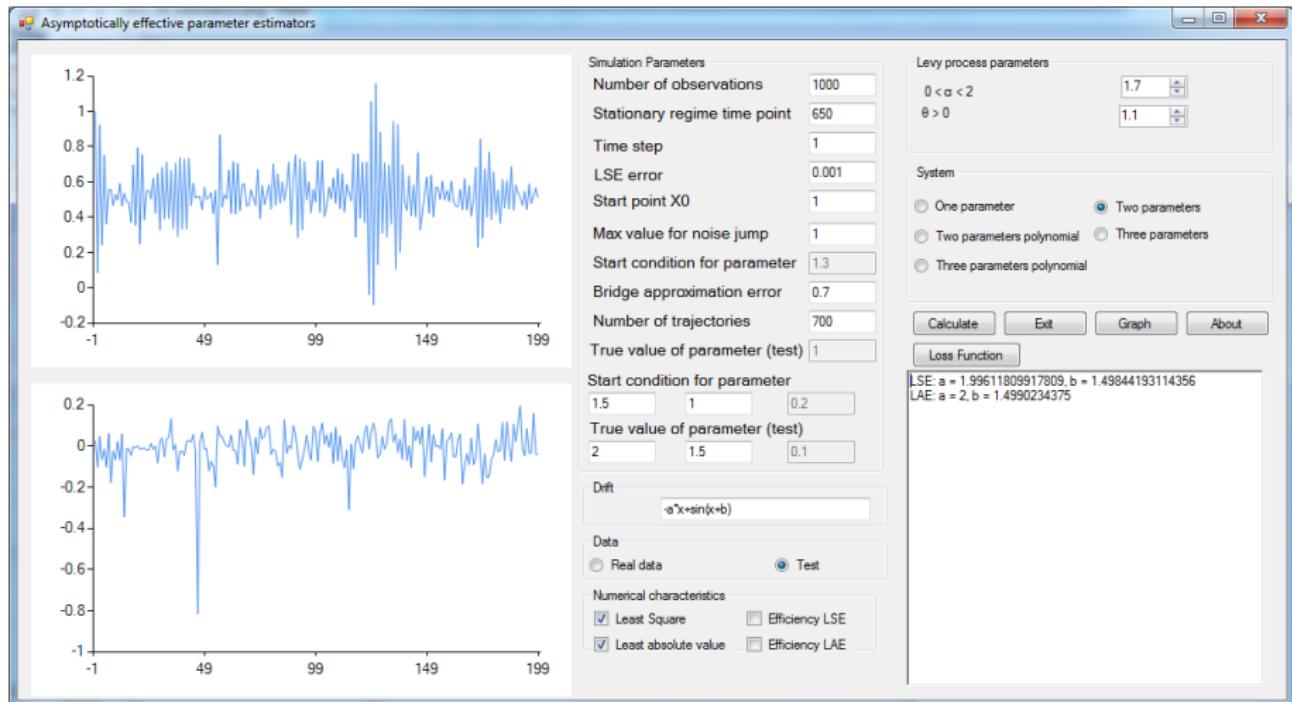


Figure: LSE and LAE in two-parameters nonlinear model.



Example 3.

Figure: Efficiency in three-parameter nonlinear model.

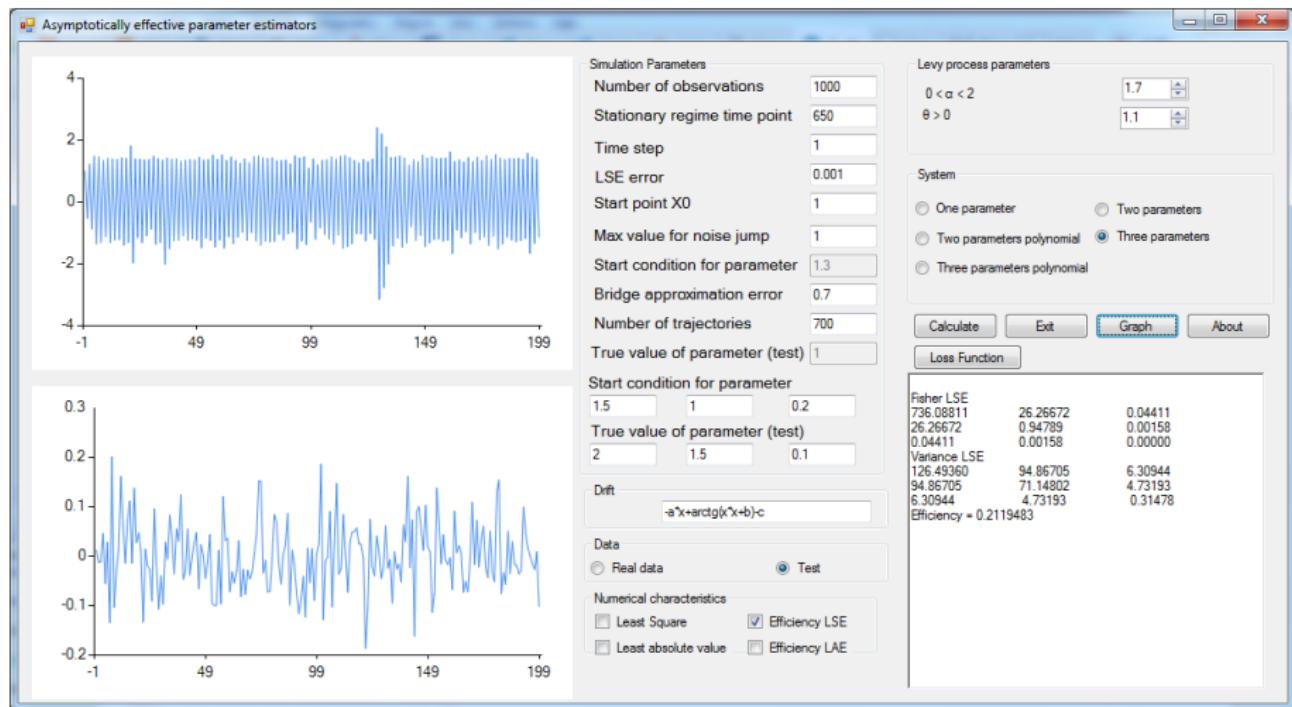
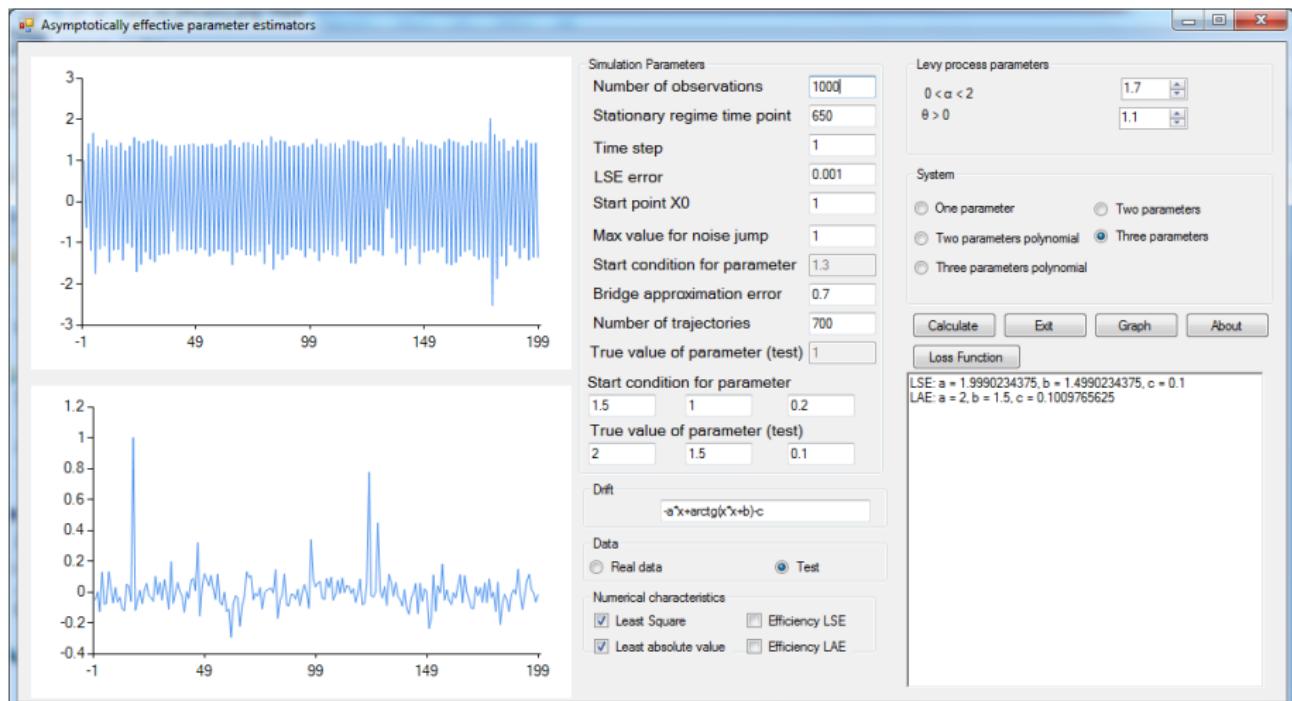


Figure: LSE and LAE in three-parameter nonlinear model.



One-Step estimator

To produce a consistent estimator with asymptotic variance equal to the inverse Fisher information as this is the best possible variance we can achieve for consistent estimators.

- Newton's Method: $\theta_0^{NM} = \theta^{LSE}$ or (like in example) θ^{LP} , $p = 3$;

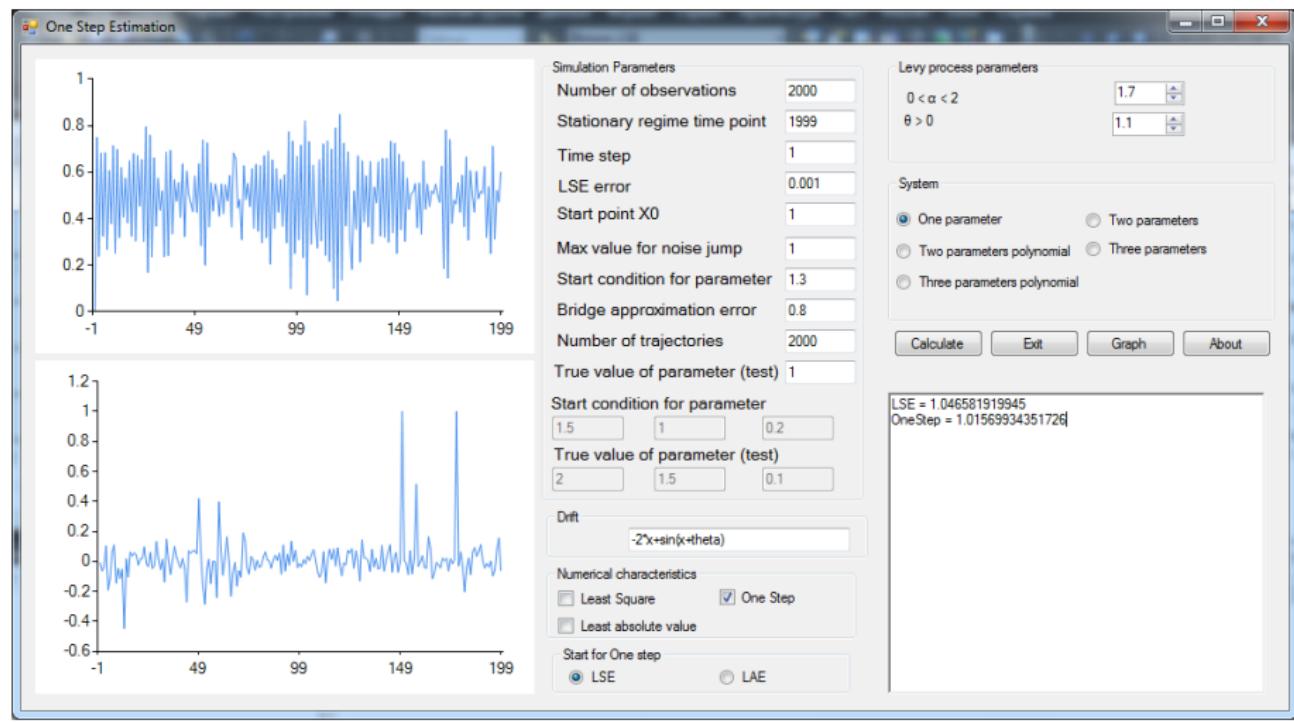
$$\theta_{k+1}^{NM} = \theta_k^{NM} - \left(\partial_{\theta\theta}^2 l(\theta_k^{NM}) \right)^{-1} \partial_\theta l(\theta_k^{NM}).$$

- Method of Scoring (Rao's accumulation method): $\theta_0^{NM} = \theta^{LSE}$,

$$\theta_{k+1}^{SM} = \theta_k^{SM} + \left(I(\theta_k^{SM}) \right)^{-1} \partial_\theta l(\theta_k^{SM}).$$

Example 4.

Figure: One-Step estimator nonlinear model.



Second part

Consider the stochastic differential equation of the form

$$dX_t = a(\alpha; X_t)dt + \sigma(\beta; X_t)dW_t,$$

where $X_0 = x_0$ and $\theta = (\alpha, \beta) \in \Theta \subseteq \mathbb{R}^2$ are unknown parameters, W is a one dimensional Wiener process and the coefficients a and σ are differentiable functions with bounded derivatives.

The following tasks are investigated:

- Develop a software to obtain a decomposition of the transition probability density in degrees t and in terms of Hermit functions.
- Use the previous decomposition to construct QMLE and One-Step estimators.

Density decomposition

We say that the function g is of order at least r if there exists some positive C and c such that

$$|g_t(\theta, x, y)| \leq Ct^r\phi_{ct}(x, y),$$

ϕ is the transition probability density at y for the Wiener process at time t which starts at x .

We want to obtain a decomposition of transition probability density of X in the form

$$p = \sum_{k,m} c_{k,m} t^k \Upsilon^{(m)}(a', b') + r,$$

where r is the function of fixed order at least n and Υ is the Hermite function given by ($m \geq 0$, H_m is the Hermite polynomial)

$$\Upsilon_t^{(m)}(a, b, x, y) = H_m(y - x - at, bt) \phi_{bt}(y - x - at), \quad b = \sigma^2.$$

The basic iterative formula of order $N/2$

$$p_t^N = p_t^0 + p_t^{N-1} \circledast \Phi, \quad p_t^0 = \gamma_t^{(0)}(a, b), \quad \Phi = \Phi_a + \Phi_b.$$

where Φ is an auxiliary kernel, \circledast is the space-time convolution.

$$\begin{aligned}\Phi_b := & \frac{1}{2}(b' - b)\gamma_t^{(2)}(a, b) = \\ & \frac{1}{2} \sum_{k=1}^N \frac{(-1)^k}{k!} (y - x)^k \gamma_t^{(2)}(a, b) \partial_y^k b + R(N/2 + \gamma/2 - 1),\end{aligned}$$

$$\begin{aligned}\Phi_a := & (a' - a)\gamma_t^{(1)}(a, b) = \\ & \sum_{k=1}^{N-1} \frac{(-1)^k}{k!} (y - x)^k \gamma_t^{(1)}(a, b) \partial_y^k a + R(N/2 + \gamma/2 - 1).\end{aligned}$$

Origin term decomposition

$$p_t^0 := \gamma_t^{(0)}(a, b) = \gamma_t^{(0)}(a', b') +$$

$$\sum_{k=1}^N \frac{t^k}{k!2^k} \gamma_t^{(2k)}(a', b') \left(\sum_{m=1}^{[N/k]} \frac{(y-x)^m}{m!} \partial_x^m b' \right)^k +$$

$$\sum_{k=1}^N \frac{1}{k!2^k} \left(\sum_{m=1}^{[N/k]} \frac{(y-x)^m}{m!} \partial_x^m a' \right)^k.$$

$$\sum_{i=1}^N \frac{t^{i+k}}{i!} \gamma_t^{(2i+k)}(a', b') \cdot \left(\sum_{j=1}^{[N/k-1]} \frac{(y-x)^j}{j!} \partial_x^j a' \right)^i +$$

$$\sum_{k=1}^N \frac{t^k}{k!} \gamma_t^{(k)}(a', b') \left(\sum_{m=1}^{[N/k-1]} \frac{(y-x)^m}{m!} \partial_x^m a' \right)^k + R(N/2 + \gamma/2).$$

Algorithm 1

1. If there is at least one difference in the syntax tree $y - x$, then go to step 2, otherwise, go to step 6.
2. Perform the transformation of the expression represented by an abstract syntax tree into a polynomial using the rules of multiplication of polynomials and using a polynomial theorem for increasing polynomials to an integer degree.
3. Simplify the abstract syntax tree.
4. While in a syntax tree it is possible to select at least one product of the Hermite function Υ on $y - x$, perform its transformation using the formula below.
5. Go to step 1.
6. Simplify the abstract syntax tree.
7. Remove terms with order greater than $N/2$. The order of the term refers to the degree of variable t .
8. End.

$$(y - x) \Upsilon_t^{(m)}(a, b) = a't \Upsilon_t^{(m)}(a, b) + bt \Upsilon_t^{(m+1)}(a, b) + m \Upsilon_t^{(m-1)}(a, b).$$

Convolution programming

$$\begin{aligned} & \int_0^t A_1(x) s^{k_1} \gamma_s^{(m_1)}(a', b') * A_2(y) (t-s)^{k_2} \gamma_{t-s}^{(m_2)}(a, b) ds = \\ &= A_1(x) A_2(y) t^{k_1+k_2+1} \left(\gamma_t^{(m)}(a', b') B(k_1 + 1, k_2 + 1) + \right. \\ & \quad \sum_{k=1}^{M1} \frac{t^k}{k! 2^k} \sum_{j=1}^{M2} \frac{t^j}{j!} \gamma_t^{(m+2k+j)}(a', b') (a' - a)^j (b' - b)^k \times \\ & \quad \left. B(k_1 + k + j + 1, k_2 + 1) + \right. \\ & \quad \sum_{k=1}^{M3} \frac{t^k}{k! 2^k} \gamma_t^{(m+2k)}(a', b') (b' - b)^k B(k_1 + k + 1, k_2 + 1) + \\ & \quad \left. \sum_{k=1}^{M4} \frac{t^k}{k!} \gamma_t^{(m+k)}(a', b') (a' - a)^k B(k_1 + k + 1, k_2 + 1) \right) + R. \end{aligned}$$

Algorithm 2

1. Let P — set of terms of the first polynomial and F — set of terms of the second polynomial.
2. Find the set of Cartesian product $S = P \times F$.
3. For each element S apply the formula above. Find the sum of all results.
4. Decompose the functions that occur as a result and depend on y in the Taylor series at the point x .
5. Perform a tree simplification that is the result of the transformations.
6. End.

Algorithm 3

1. Using **Algorithm 1** rewrite in terms of Hermite functions a zero term p_t^0 and auxiliary kernel Φ .
2. $i := 1$.
3. Using **Algorithm 2** to find polynomial $p_t'^i = p_t^{i-1} \otimes \Phi$.
4. Using **Algorithm 1** rewrite $p_t'^i$ in terms of Hermite functions.
5. Find $p_t^i = p_t^0 + p_t'^i$.
6. Do simplification p_t^i .
7. Remove from polynomial p_t^i terms with an order greater than $N/2$.
8. If $i < N$, then go to step 9, otherwise, go to step 10.
9. $i := i + 1$.
10. End.

Programm interface

Calculator

Order: 2

Generate

a(x) := sin(x)

b(x) := cos(x)

Substitute

Variables one per line with syntax 'variableName=expression' where expression should evaluate to numeric value and can depend on already declared variables

x=1.0
t=x*PI

Calculate

Result:

Export as LaTeX

About

Guide

Result of execution

$$\begin{aligned} p_t^2 = & \gamma_t^{(0)}(a, b) + \frac{1}{4} \cdot b \cdot \partial_x(b) \cdot t^2 \cdot \gamma_t^{(3)}(a, b) + \\ & \left(\frac{1}{2} \cdot b \cdot \partial_x(a) + \frac{1}{8} \cdot b \cdot \partial_x^2(b) \right) \cdot t^2 \cdot \gamma_t^{(2)}(a, b) + \\ & \left(\frac{1}{4} \cdot a \cdot \partial_x(b) + \frac{3}{2} \cdot (\partial_x(b))^2 \right) \cdot t^2 \cdot \gamma_t^{(2)}(a, b) + \\ & \left(\frac{5}{4} \cdot b \cdot (\partial_x(b))^2 + \frac{1}{12} \cdot \partial_x^2(b) \cdot (b)^2 \right) \cdot t^3 \cdot \gamma_t^{(4)}(a, b) + \\ & \frac{5}{32} \cdot (\partial_x(b))^2 \cdot (b)^2 \cdot t^4 \cdot \gamma_t^{(6)}(a, b). \end{aligned}$$

QMLE and LSE

Diffusion

Parameters

Interval	10000	True alpha	1
Error for optimisation	0.001	True beta	0.5
Step	0.8	Start alpha	1.5
Start point	1	Start beta	0.1

Stochastic equation

$$dX = -\alpha * x \, dt + 2 * \sin(x * \beta) \, dW$$

QMLE: alpha = 0.999585725112595, beta = 0.506658337193061
LP: p = 2, 0.999298402980439, beta = 0.635254914051872

Optimisation parameter

Proportion coefficient	0.04
Variation interval	0.7

Estimation

LP Cov LSE
 QMLE Cov QMLE
 Asymptotic variance

LP Power 2

Efficiency

N of trajectories 700

Graph About Calculate Exit



Asymptotic properties of estimators of unknown SDE's parameters Potsdam, 25.03 – 03.04.2019 30 / 33

QMLE and LAE

Diffusion

Parameters

Interval	10000	True alpha	1.3
Error for optimisation	0.001	True beta	0.9
Step	0.8	Start alpha	1.5
Start point	1	Start beta	0.8

Stochastic equation

$$dX = -\alpha * x \, dt + 2 + \arctg(x^* x + t) \, dW$$

QMLE: alpha = 1.282166202031746, beta = 0.934193991392144
LP: p = 1, alpha = 1.374114647944737, beta = 0.835700657715283

Optimisation parameter

Proportion coefficient	0.04
Variation interval	0.7

Graph About Calculate Exit

Estimation

Lp Cov LSE
 QMLE Cov QMLE
 Asymptotic variance

LP Power 1

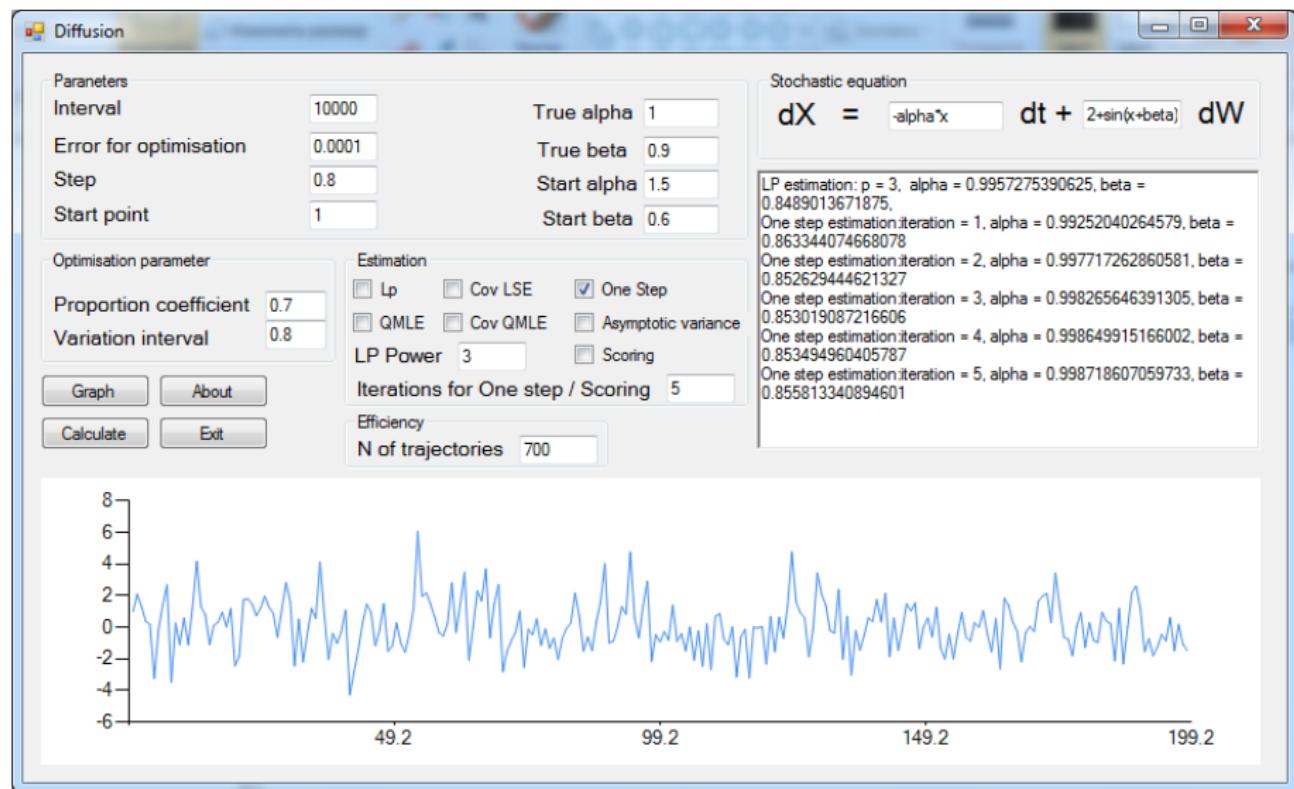
Efficiency

N of trajectories 700



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One-Step estimator (Newton's method)



One-Step estimator (Rao's accumulation method)

